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Pulsation analysis and its impact on primary transit modeling in WASP-33*,**

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Received 6 August 2013 / Accepted 19 October 2013

ABSTRACT

Aims. To date, WASP-33 is the only δ Scuti star known to be orbited by a hot Jupiter. The pronounced stellar pulsations, showing periods comparable to the primary transit duration, interfere with the transit modeling. Therefore our main goal is to study the pulsation spectrum of the host star to redetermine the orbital parameters of the system by means of pulsation-cleaned primary transit light curves.

Methods. Between August 2010 and October 2012 we obtained 457 h of photometry of WASP-33 using small and middle-class telescopes located mostly in Spain and in Germany. Our observations comprise the wavelength range between the blue and the red, and provide full phase coverage of the planetary orbit. After a careful detrend, we focus our pulsation studies in the high frequency regime, where the pulsations that mostly deform the primary transit exist.

Results. The data allow us to identify, for the first time in the system, eight significant pulsation frequencies. The pulsations are likely associated with low-order p-modes. Furthermore, we find that pulsation phases evolve in time. We use our knowledge of the pulsations to clean the primary transit light curves and carry out an improved transit modeling. Surprisingly, taking into account the pulsations in the modeling has little influence on the derived orbital parameters. However, the uncertainties in the best-fit parameters decrease. Additionally, we find indications for a possible dependence between wavelength and transit depth, but only with marginal significance. A clear pulsation solution, in combination with an accurate orbital period, allows us to extend our studies and search for star-planet interactions (SPI). Although we find no conclusive evidence of SPI, we believe that the pulsation nature of the host star and the proximity between members make WASP-33 a promising system for further SPI studies.

Key words. asteroseismology – instrumentation: photometers – planet-star interactions – methods: observational – techniques: photometric – stars: variables: δ Scuti

1. Introduction

The first mention of δ Scuti variability was made more than one hundred years ago (Campbell & Wright 1900). Fifty years later, Eggen (1956) pointed out the need to place these variable stars under an independent stellar classification. The δ Scutis have been among us for a long time. Nevertheless, because of the intrinsically small variability near the limit of detectability major studies of them did not begin until the 1970s (Baglin et al. 1973; Breger 1979; Breger & Stockenhuber 1983). Nowadays, the *Kepler* telescope alone provides precise light curves of several hundred δ Scuti stars (e.g., Uytterhoeven et al. 2011).

In the Hertzsprung-Russell diagram, the δ Scuti stars are located in an instability strip covering spectral types between A and F (Baglin et al. 1973; Breger & Stockenhuber 1983). Most

 δ Scuti stars belong to Population I (Breger 1979), with typical masses of 2 M_{\odot} (Milligan & Carson 1992). Due to both radial and nonradial pulsations, δ Scuti stars show brightness variations from milli-magnitudes up to almost one magnitude in blue bands. In δ Scuti stars, pulsations are driven by opacity variations. There are two distinct types of pulsation modes that might occur (e.g., Breger et al. 2012): short-period p-modes (pressure modes, for which pressure serves as the restoring force) and long-period g-modes (gravity modes, with buoyancy as the restoring force). A typical δ Scuti pulsation spectrum shows dozens of periods (Breger et al. 1999a,b), with cycle durations ranging from a couple of hours to the minute regime.

WASP-33 (HD 15082) is a bright ($V \sim 8.3$), rapidly rotating ($v \sin(i) \sim 90 \text{ km s}^{-1}$)¹ δ Scuti star; in fact, it is both the hottest and only δ Scuti star known to date to host a hot Jupiter (Christian et al. 2006). The planet, WASP-33b, was detected through its transits in the frame of the WASP campaign (Pollacco et al. 2006). It circles its host star every 1.22 d in

^{*} Tables 1 and 10 and Fig. 8 are available in electronic form at http://www.aanda.org

^{**} Photometry is only available at the CDS via anonymous ftp to cdsarc.u-strasbg.fr (130.79.128.5) or via http://cdsarc.u-strasbg.fr/viz-bin/qcat?J/A+A/561/A48

 $^{^{1}}$ i corresponds to the inclination of the stellar rotation axis.

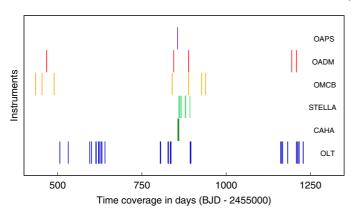


Fig. 1. Sampling of our observations.

a retrograde orbit. With a brightness temperature of 3620 K, WASP-33b is the hottest exoplanet known to date (Smith et al. 2011). Showing an unusually large radius, WASP-33b belongs to the class of anomalously inflated exoplanets (Collier Cameron et al. 2010). For its mass and hence density, only an upper limit of $M \sin(i) < 4.59 \ M_{\rm J}^2$ has been determined.

The host star, WASP-33A, shows pronounced pulsations with periods on the order of one hour. Collier Cameron et al. (2010) note that the presence of these pulsations offers "the intriguing possibility that tides raised by the close-in planet may excite or amplify the pulsations in such stars". The discovery of WASP-33's pulsations within photometric data were first reported by Herrero et al. (2011), who suggest a possible commensurability between a pulsation period and the planetary orbital period with a factor of 26, indicative of SPI.

We study the pulsations and primary transits using a total of 56 light curves of WASP-33, observed during two years and providing complete orbital phase coverage.

2. Observations and data reduction

Between Aug. 2010 and Oct. 2012, we obtained 457 h of photometry of WASP-33 distributed across 56 nights using six telescopes: one in Germany and five in Spain. Figure 1 shows the temporal coverage provided by the individual telescopes; the details of the observations are given in Tables 1 and 2 and the technical characteristics of the telescopes are summarized in Table 3.

Throughout the analysis, the barycentric dynamical time system is used (BJD_{TDB}). Conversions between different time reference systems have been carried out using the web-tool made available by Eastman et al. (2010)³.

WASP-33 is located in a sparse stellar field. Therefore the defocusing technique did not produce any undesired effect, such as overlapping of the stellar point spread functions. However, after either defocusing or considering the natural seeing of the sites, the optical companion identified by Moya et al. (2011) is contained, in most of the cases, inside the selected aperture radius. A discussion on third-light contribution is addressed in Sect. 4.

2.1. Hamburger Sternwarte

The Oscar Lühning Telescope⁴ (OLT) is located at the Hamburger Sternwarte, Germany. It is equipped with an Apogee

Table 2. Overview of observation nights for CAHA, OADM, OMNP, and OMCB.

Date	Duration (h)	F^a	NoP ^b	Airmass			
САНА							
2011 October 22	4.1	v, b, y	130	$1.0 \to 1.9$			
2011 October 23	3.2	v, b, y, d	178	$1.3 \rightarrow 2.2$			
2011 October 24	3.4	v, b, y	157	$1.0 \rightarrow 1.9$			
2011 October 25	11.9	v, b, y	462	$1.0 \rightarrow 2.2$			
OADM							
2010 September 28	3.7	R	279	$1.0 \to 1.3$			
2011 October 10	3.5	V	344	$1.0 \rightarrow 1.5$			
2011 November 23	4	V	320	$1.0 \rightarrow 1.1$			
2012 September 24	4.2	V	630	$1.0 \rightarrow 1.0$			
2012 October 8	4	V	588	$1.0 \rightarrow 1.1$			
OMNP							
2011 October 21	4.6	V	439	$1.0 \rightarrow 1.5$			
OMCB							
2010 August 26	6.4	R	162	$1.0 \rightarrow 2.2$			
2010 September 14	5.3	R	125	$1.0 \rightarrow 1.2$			
2010 October 20	5.6	R	138	$1.0 \rightarrow 1.7$			
2011 October 5	7.6	R	195	$1.0 \rightarrow 1.4$			
2011 November 23	5.1	R	127	$1.0 \rightarrow 1.3$			
2012 January 1	4.5	V	112	$1.0 \rightarrow 1.2$			
2012 January 12	5.2	V	130	$1.0 \rightarrow 1.4$			

Notes. (a) Filter; (b) Number of photometric data points.

Table 3. Technical telescope data: primary mirror diameter, field of view (FOV), plate scale, and observatory location.

Name		FOV ^a	Scale	Locationa
	(m)		("/pix)	
OLT	1.2	9' × 9'	0.158	G
CAHA	2.2	$18' \times 18'$	0.135	S
STELLA	1.2	$22' \times 22'$	0.322	CI
OADM	0.8	$12' \times 12'$	0.36	S
OMNP	0.4	$21' \times 21'$	1.24	S
OMCB	0.3	$16' \times 11'$	0.62	S

Notes. ^(a) Full fields of view are listed. ^(b) Germany (G), continental Spain (S), Canary Islands (CI).

Alta U9000 charge-coupled device (CCD) camera with guiding system.

Between Nov. 2010 and Oct. 2012, we observed WASP-33 for 29 nights using the Johnson-Cousins B and R filters (see Table 1 for details). The exposure time was between 10 and 40 s, mainly depending on the night quality. The airmass ranged from principally 1 to 3 only when photometric nights allowed such observations. The typical seeing at the Hamburger Sternwarte is 2.5-3 arcsec. Therefore saturation was not an issue. During a total time of ~ 150 h, we obtained 18 090 photometric data points providing both in- and out-of-transit coverage. Additionally, calibration frames were obtained for each individual night.

For bias subtraction and flat fielding, we used the *ccdproc* package in IRAF; aperture photometry was carried out using IRAF's *apphot*. To obtain differential photometry, we measured unweighted fluxes in WASP-33 and two reference stars using various aperture radii. The final light curve was produced using the aperture that minimizes the scatter of the differential light curve. It is based on the brighter of the two reference stars,

i corresponds to the inclination of the planetary orbit.

http://astroutils.astronomy.ohio-state.edu/time/

http:www.hs.uni-hamburg.de/

BD+36 488, which is, however, still one magnitude fainter than WASP-33. The remaining reference star was used to obtain control light curves to ensure that the photometry is based on a proper reference.

2.1.1. Calar Alto observatory

The German-Spanish Astronomical Center at Calar Alto (CAHA) is located close to Almería, Spain. It is a collaboration between the Max-Planck-Institut für Astronomie (MPIA) in Heidelberg, Germany, and the Instituto de Astrofísica de Andalucía (CSIC) in Granada, Spain.

We used the Bonn University Simultaneous CAmera (BUSCA)⁵ instrument, which is mounted at the 2.2 m telescope. The instrument allows simultaneous measurements in four different spectral bands. To reduce read-out time, we exposed only the half-central part of the CCD. We observed WASP-33 using the Strömgren v, b, and y filters, along with a filter labeled d, which is centered at 753 nm with a FWHM of 30 nm. As BUSCA requires simultaneous read-out of all CCDs, we used an exposure time of 4 s, which provides adequate signal-to-noise ratios in the Strömgren bands and avoids saturation of the source. The photometry obtained by means of the d filter was discarded due to low signal-to-noise. During our observations, the seeing ranged between 1 and 1.5 arcsec. In the visible, the extinction was between 0.15 and 0.2 mag/airmass. Observing for ~23 h, we obtained 927 photometric data points per filter. The data reduction was carried out as described in Sect. 2.1.

2.2. Observatorio del Teide

STELLar Activity (STELLA)⁶ consists of two fully robotic 1.2 m telescopes, one dedicated to photometry and the other to spectroscopy (Strassmeier et al. 2010). The photometer is a wide-field imager called WiFSIP. It is equipped with a 4092^2 15- μ m pixel back-illuminated CCD.

We observed WASP-33 using STELLA for one month starting at the end of Oct. 2011. STELLA's optical setup offers a field of view of $22' \times 22'$. However, for the purposes of these observations and with the main goal of reducing readout times, we used only a $15' \times 15'$ subframe. To obtain quasi-simultaneous multiband photometry, we alternated between the Strömgren v and b filters. In this way, we obtained 2483 photometric measurements with the v filter and 2234 with the v filter, which equates to v05 h per spectral range. Accounting for read-out time, the typical temporal cadence was of v00 s. The optics had to be defocused to avoid saturation of the target.

We carried out the data reduction using ESO-MIDAS. Bias frames were obtained every night, evening, and morning and combined into a master bias on a daily basis. Twilight flat-field frames for both the Strömgren b and v filters were obtained approximately every 10 d. Bias subtraction and flat-fielding were performed as usual. We carried out aperture photometry using SExtractor's MAG AUTO option. Here, SExtractor computes an elliptical aperture for every detected object in the field, following its light distribution in x and y, and scales the aperture width with the SExtractor parameter k, which we set to 2.6. As a flux calibrator, we used the summed, unweighted flux of three reference stars, viz., BD+36 493, BD+36 487, and BD+36 488. The background, estimated locally for each object, was generally low.

2.3. Primary transit observations

To increase our sample of primary transit light curves, we used three telescopes with apertures between 0.3 and 0.8 m. To carry out the observations we defocused the telescopes. In the particular case of bright sources such as WASP-33, long exposures after defocusing can reduce scintillation noise and flat-fielding errors (e.g., Southworth et al. 2009; Gillon et al. 2009). In this way, we reached milli-magnitude precision in all of the primary transit light curves obtained using small-aperture telescopes.

The Telescopi *Joan Oró* is a fully robotic telescope located at the Observatori Astronòmic del Montsec⁷ (OADM). It is equipped with an FLI Proline 4240 CCD and standard Johnson-Cousins filters. Observing for five nights distributed over two years, we collected 19.4 h of data at a temporal cadence of \sim 45 s.

The Observatori Montcabrer⁸ (OMCB) is located in Cabrils, Spain. We used its remotely operated 0.3 m telescope, which is equipped with an SBIG ST-8 CCD and standard Johnson-Cousins filters for seven nights distributed over 1.5 yr. In total, we collected ~40 h of data with typical exposure times of ~120 s. Although the observatory is located in a light-polluted area, a photometric precision of about 1 mmag could be reached.

The Observatori Món Natura Pirineus⁹ (OMNP) contributed a 0.4 m telescope equipped with an SBIG STL-1001E CCD. With this telescope we observed one primary transit using the Johnson-Cousins *V* filter.

The data of all the telescopes were corrected for bias and dark current and were flat-fielded using MaximDL and new calibration images for every night. The light curves were produced using Fotodif¹⁰. In all cases, the aperture radius was selected such that the scatter in the out-of-transit sections of each light curve is minimized.

3. WASP-33 as a δ Scuti star

The primary transit light curves of WASP-33 are deformed by the host star's pulsations. This interferes with transit modeling and therefore with the determination of the orbital and physical parameters of the system. As the removal of an inappropriate primary transit model could introduce a spurious signal in the pulsation spectrum of the star, associated with the planetary orbital period rather than intrinsic stellar variability, we use only off-transit data points to determine WASP-33A's pulsation spectrum.

The pulsation frequency analysis is performed using PERIOD04, a package intended for the statistical analysis of large astronomical data sets containing gaps, with single-frequency and multiple-frequency techniques (Lenz & Breger 2005). The package utilizes both Fourier and multiple-least-squares algorithms, which do not rely on sequential prewhitening or assumptions of white noise.

3.1. Light curve normalization

To avoid variations apart from the periodicity that we want to characterize, the light curves need to be first detrended. Thus, to study the pulsation spectrum in the high-frequency regime, we normalize each individual light curve. In this way, we eliminate

http://www.caha.es/CAHA/Instruments/BUSCA/intro.html

⁶ http://www.aip.de/stella/

http://www.oadm.cat/en/

⁸ http://cometas.sytes.net/

http://monnaturapirineus.com/en/content/observatory

¹⁰ http://www.astrosurf.com/orodeno/fotodif/index.htm

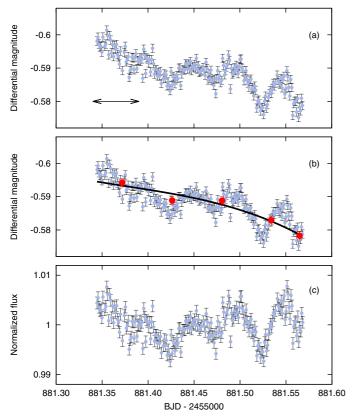


Fig. 2. Our normalization procedure: **a)** Differential light curve in magnitudes, obtained using the STELLA telescope and the Strömgren v filter. The length of the arrow indicates the width of the time-bins. **b)** Thick red points: the binned time, flux, and photometric error. Continuous black line: third-degree polynomial fitted to the binned data points. **c)** Normalized light curve in flux units.

the low-frequency signals that might be associated with systematic effects, such as residual fluctuations due to atmospheric extinction, that are unrelated to intrinsic stellar variability.

Our procedure is the following: first, we bin the light curves using time-bins with a duration of ~ 1.3 h to "hide" the high-frequency pulsations inside them. Second, we calculate the mean value of time and flux in the bins and fit a low-order polynomial to the binned light curves. The degree of the polynomial depends on the number of available data points and therefore on the duration of the observing night. Finally, we subtract the fitted polynomial from the unbinned light curve and convert magnitudes into flux. To ensure a proper normalization, we visually inspected the results of our procedure. Figure 2 shows a representative example.

3.2. Light curve normalization and its relevance for periodogram analysis

To study the impact of the normalization on the high-frequency domain, we applied an alternative normalization and subtracted only the mean value of the light curves. We then compared the periodograms obtained for each data set based on the two normalizations. Figure 3 shows the resulting power spectra for our STELLA data obtained with the Strömgren b filter as an example. The vertical dashed line in the figure indicates the frequency corresponding to the average length of our observing nights. The difference between the periodograms shows how the normalization process affects the power spectrum. The discrepancy is

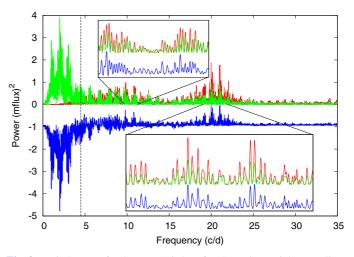


Fig. 3. Periodogram for STELLA b data for the polynomial normalization in red, the alternative normalization with a constant in green, and their difference in blue (arbitrarily shifted). The vertical dashed line indicates the mean duration of the observing nights.

strongest for frequencies corresponding to periods longer than the average night length. Beyond that limit, the effect of the normalization becomes weak. Figure 3 shows close-ups of the periodograms around $\nu \sim 20$ c/d and $\nu \sim 10$ c/d. While the amplitudes of individual peaks change, the structure of the periodogram and the position of the peaks remain stable.

To verify the stability of the peak positions, we searched both periodograms for strong peaks and compared their frequency. To determine the peak positions, we used a Gaussian fit. Based on 66 peak pairs in the >10 c/d regime, we derive a shift of -0.003 ± 0.0025 c/d. At lower frequencies the uncertainty becomes larger, which is in agreement with the behavior observed in Fig. 3. Thus, we conclude that the normalization does not seriously impede the analysis in the high-frequency regime.

3.3. Determination of WASP-33's frequency spectrum

Our frequency analysis is based on the STELLA, CAHA, and OLT data. The latter provide a total temporal coverage of approximately two years, essentially concentrated, however, in three observing seasons. Each OLT season was considered separately in our frequency search. We were left with five data sets obtained in four different spectral filters. Although the photometric amplitudes of δ Scuti stars depend on wavelength, it is possible to combine multifilter data to determine the frequencies via Fourier methods. Consequently, we combine the data obtained with the v, B, and b filters, for which we found the difference in amplitude values to be statistically insignificant.

We identified the frequencies of the dominating pulsations by analyzing the combined data set, which provides the cleanest spectral window and the highest precision in the determined frequencies.

To estimate the signal-to-noise level of a given pulsation with amplitude $A_{\rm o}$, we computed the average amplitude, $\sigma_{\rm res}$, over a frequency interval with a width of 2 c/d from a periodogram obtained from the final residuals (see Fig. 4) and estimated the amplitude signal-to-noise ratio (ASNR) of each pulsation as $A_{\rm o}/\sigma_{\rm res}$. Following Breger et al. (1993), we consider a pulsation to be significant when the estimated ASNR of the periodogram peak is larger than four.

The residual power spectrum indicates strong departures from white noise arising from a potentially highly complex

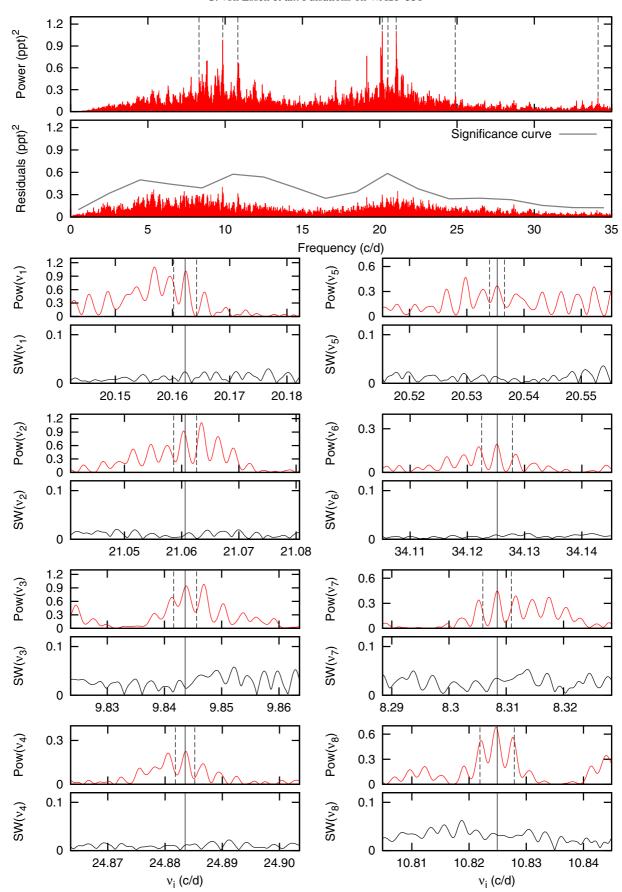


Fig. 4. *Top panels*: periodogram of the combined data set and the final residuals in part per thousand (ppt). The solid line indicates the significance curve. *Bottom panels*: a closer look at the periodogram around the detected frequencies and the spectral window (SW). Dashed vertical lines around the peaks indicate our error estimates as listed in Table 4.

Table 4. Parameters of the pulsations with $1-\sigma$ errors. Frequency ratios (FR) of the pulsations and the orbit (see Sect. 5.3).

PN	$v \pm \sigma_v (c/d)$	Combined data $A \pm \sigma_A (10^{-3})$	$\phi \pm \sigma_{\phi} (2\pi)$	ASNR	Subsample analysis $v \pm \sigma_v$ (c/d)	FR
Puls ₁	20.16214 ± 0.00063	0.95 ± 0.03	0.5718 ± 0.0049	7.8	20.1621 ± 0.0023	24.595
Puls ₂	21.06057 ± 0.00058	0.93 ± 0.03	0.3594 ± 0.0050	7.6	21.0606 ± 0.0023	25.691
Puls ₃	9.84361 ± 0.00066	0.79 ± 0.03	0.4804 ± 0.0058	6.3	9.8436 ± 0.0023	12.008
Puls ₄	24.88351 ± 0.00056	0.42 ± 0.03	0.3009 ± 0.0110	4.3	24.8835 ± 0.0017	30.355
Puls ₅	20.53534 ± 0.00057	0.71 ± 0.03	0.5207 ± 0.0065	5.9	20.5353 ± 0.0013	25.050
Puls ₆	34.12521 ± 0.00054	0.49 ± 0.03	0.5738 ± 0.0096	8.0	34.1252 ± 0.0027	41.628
Puls ₇	8.30842 ± 0.00054	0.63 ± 0.03	0.5173 ± 0.0074	4.2	8.3084 ± 0.0025	10.135
$Puls_8$	10.82492 ± 0.00058	0.64 ± 0.03	0.9969 ± 0.0072	5.8	10.8249 ± 0.0030	13.205

pulsation spectrum and/or aliasing problems. Nonetheless, all remaining peaks remain below our significance curve.

Because we have made annual solutions, the phase and amplitude shifts from year to year have been taken care of. However, if small and systematic changes occur from year to year, there will exist smaller changes within each observing season. Such changes are not taken care of by PERIOD04 and will lead to close side lobes in the periodogram. However, since the solution presented in this work does not contain very close frequencies, these side lobes will not affect it.

Strong aliasing represents an unavoidable difficulty, leading to 1 c/d ambiguities and a large number of strong peaks at the 30% level (relative to the main peak). This aliasing, combined with a large number of frequencies, sets the limits of our multifrequency analysis. To minimize the effect of aliasing and the window function on our frequency selection, we checked that the significant peaks are present in all subsamples, where we, however, tolerated lower ASNR levels. As the subsamples comprise only a fraction of the data, the frequencies cannot be determined with the same accuracy as in the combined sample. We obtained estimates of the frequency and the uncertainty (see Sect. 3.4) from the subsamples and verified that the results are consistent with the values obtained using the combined data set. Only if this was the case did we accept a frequency. One frequency near 7.3 c/d, detected in the combined data set, could not be found in all individual data sets and is consequently not included in our analysis.

Altogether, eight frequencies were extracted from the data. Table 4 shows the frequencies, amplitudes, phases, and associated ASNRs. Additionally, we provide the mean frequency and error obtained from the subsamples. Figure 5 shows our pulsation model, plotted over some of the available off-transit light curves. The displayed pulsation model is obtained after fitting the phases to each individual night only. The reasons for this procedure will be given in Sect. 3.5.

3.3.1. Analysis of fit quality

To quantify the improvement in the description provided by our pulsation model, we calculate the resulting χ^2 values for the pulsation model and a constant, $\chi^2_{\mathbb{C}}$ and χ^2_{mod} and carry out an F-test. In particular, we calculate the F-statistics using

$$F = \frac{(\chi_{\text{mod}}^2 - \chi_{\mathbb{C}}^2)/(\nu_{\text{mod}} - \nu_{\mathbb{C}})}{\chi_{\mathbb{C}}^2/\nu_{\mathbb{C}}},\tag{1}$$

where ν corresponds to the degrees of freedom, $\nu_{\rm mod} = 14\,701$, and $\nu_{\rm C} = 14\,725$. Formally, we obtain a p-value of 1×10^{-16} , indicating that the pulsation model accounts for a substantial fraction of the light curve variations.

Although our model reproduces the overall stellar pulsation pattern, the bottom panels of Fig. 5 show flux residuals that do not behave as random, uncorrelated noise. Such residuals may be produced by nonsinusoidal pulsations, low-amplitude pulsations not accounted for in the model or by local changes in the atmospheric conditions not entirely removed by the differential photometry. At any rate, the remaining scatter in the data defines the limiting accuracy achievable in cleaning the primary transits.

3.4. Error treatment

Preliminary error estimates for the frequencies listed in the second column of Table 4 were obtained in two ways. First, we followed the analytical expressions of Breger et al. (1999a). Second, we fitted a Gaussian function to the peaks and used the standard deviation as our error estimate. To be conservative, we used the larger of these as frequency-error estimate. As the residuals show correlated noise, the true uncertainties in our frequencies could be considerably larger. A deeper discussion of errors is given below.

3.4.1. Correlated noise and unevenly spaced data for periodogram analysis

Montgomery & O'Donoghue (1999) present analytical results of the effect that random, uncorrelated noise has on a least-squares fit of a sinusoidal, evenly sampled signal. They provide the following expressions for the uncertainties:

$$\sigma_{\nu} = \frac{\sqrt{6} \,\sigma_{N}}{\pi \,\sqrt{N}AT} \tag{2}$$

$$\sigma_A = \sqrt{\frac{2}{N}} \sigma_N \tag{3}$$

$$\sigma_{\phi} = \frac{\sigma_N}{\pi \sqrt{2N}A},\tag{4}$$

where σ_{ν} , σ_{A} , and σ_{ϕ} are the standard deviations for a sinusoidal signal with frequency ν , amplitude A, and phase ϕ . The remaining parameters are N for the total number of data points, T for the total duration of the observing campaign, and σ_{N} for the average measurement error of the data points. If the time series are unevenly sampled and show correlated noise (e.g., due to atmospheric extinction), Montgomery & O'Donoghue (1999) suggest to estimate errors according to

$$\sigma^{2}(\omega) = \sigma_{0}^{2}(\omega)A(\omega, D) \tag{5}$$

$$A(\omega, D) = D \frac{\sqrt{\pi}}{2} e^{-(\frac{D\Delta t\omega}{4})^2},\tag{6}$$

where σ_0^2 is the variance of the parameter for uncorrelated data sets, given by Eqs. (2)–(4). Further, Δt is the mean exposure time

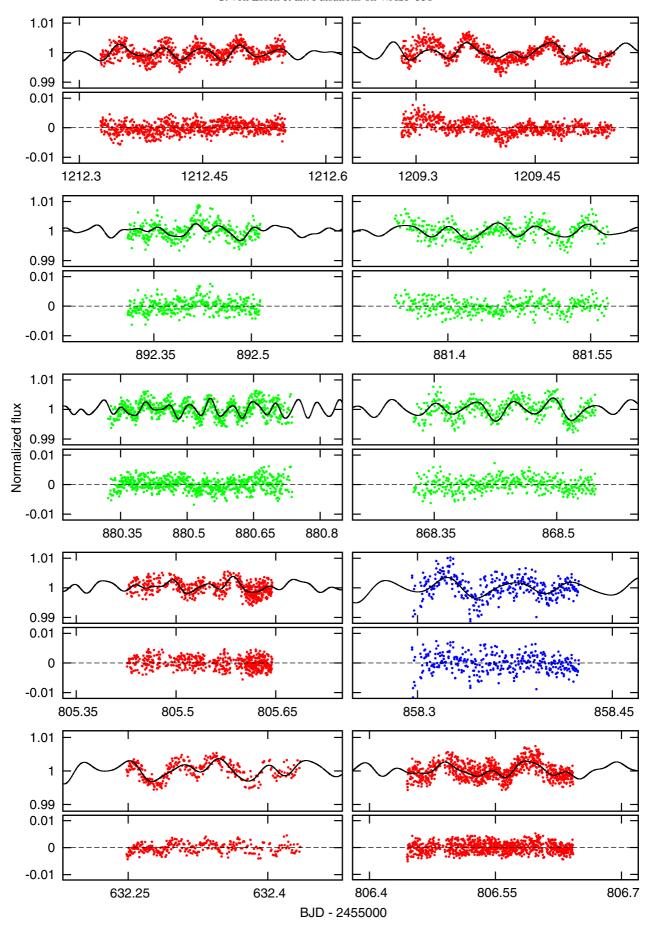


Fig. 5. Exemplary off-transit light curves color-coded in red for OLT, green for STELLA, and blue for CAHA, overplotted with the pulsation model in black continuous line. *Top panels*: normalized flux. *Bottom panels*: residuals after the pulsation model has been subtracted.

Table 5. Upper limits for the errors of the pulsation-associated parameters.

	$\langle \sigma_{\scriptscriptstyle \mathcal{V}} angle$	$\langle \sigma_A \rangle$	$\langle \sigma_{\phi} \rangle$
Puls ₁	0.009	0.18	0.031
$Puls_2$	0.008	0.18	0.031
$Puls_3$	0.021	0.38	0.078
$Puls_4$	0.007	0.15	0.058
Puls ₅	0.008	0.18	0.041
Puls ₆	0.004	0.11	0.036
Puls ₇	0.020	0.49	0.116
$Puls_8$	0.016	0.35	0.087

of the data set, D is an estimate of the number of consecutive correlated data points, and $\omega = 2\pi v$ is the angular frequency of the pulsation.

An upper limit for the error can be obtained by maximizing Eq. (6). This occurs when the correlation time is on the order of the signal period. In this case, we obtain $A = A_{\rm max} \sim 0.24 P/\Delta t$. Table 5 shows the upper-limit uncertainty estimates for our pulsation model parameters, obtained by means of Eq. (5). Here, we used the last column of Table 4 as an estimate for the "uncorrelated" errors. The estimated upper error limits remain satisfactory to characterize the pulsations photometrically.

3.4.2. Photometric errors

The photometric reduction tasks used in this work neglect systematic effects and provide statistical measurement errors, which are rather lower limits to the true uncertainties. We study the impact of the measurement errors on our frequencies analysis, based on the OLT, STELLA, and CAHA data (Johnson-Cousins *B* filter for OLT and Strömgren *v* for STELLA and CAHA).

In particular, we randomly increase the photometric errors by a factor of up to two and recalculate the position of the leading peaks and their respective ASNR. After repeating this procedure 10^4 times, we analyzed the resulting statistics of peak positions and ASNR. We find that the observed change in frequency is contained within the previously derived error. The ASNR decreases but remains higher than ~ 4 in all cases. Therefore we conclude that our frequency analysis is robust against a moderate increase of up to 100% in the photometric error.

3.5. Phase-shift analysis

Photometry provided by the *Kepler* satellite has widely been used to study the pulsation spectrum and its evolution in δ Scuti stars (e.g., Balona et al. 2012b,a; Southworth et al. 2011; Murphy et al. 2012). Most of the analyzed δ Scuti stars pulsate in several modes. For example, Breger et al. (2012) identify 349 frequencies in the rapidly rotating Sct/Dor star KIC 8054146, for which the authors even find variations in amplitude and phase.

Following the method of Breger (2005) to identify amplitude variations and phase shifts, we divided our off-transit data sets into four subsets: from BJD $\sim 2\,455\,596$ to BJD $\sim 2\,455\,641$ ($\sim 1.5\,$ months), from BJD $\sim 2\,455\,805$ to BJD $\sim 2\,455\,837$ ($\sim 1\,$ month), from BJD $\sim 2\,455\,858$ to BJD $\sim 2\,455\,896$ ($\sim 1\,$ month), and from BJD $\sim 2\,456\,162$ to BJD $\sim 2\,456\,217$ ($\sim 2\,$ months). This particular choice avoids including data gaps due to seasonal effects in the subsamples and therefore limits the impact of aliasing. As the amplitude of WASP-33's pulsations is too low to identify amplitude

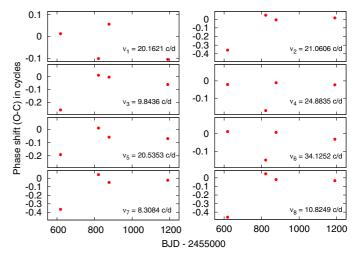


Fig. 6. Temporal phase evolution of the pulsation frequencies.

variations by means of our photometric data, we focus on phase shifts. In particular, we fit the phases in each subsample, fixing the amplitudes and frequencies to the values listed in Table 4. Figure 6 shows our results. Error bars are on the order of ~ 0.005 and therefore rather negligible in the plot.

For all eight detected pulsations, we find a change in phase. There are striking similarities between the O–C diagrams of ν_2 , ν_3 , ν_5 , ν_7 , and ν_8 , as well as between ν_4 and ν_6 . The largest observed gradient is about 2×10^{-3} d⁻¹, assuming a linear evolution. Clearly, such shifts must be taken into account in the construction of a pulsation model to clean the transits.

3.6. Mode identification

A particularly interesting question related to the observed frequencies is their association with specific pulsation modes, i.e., the radial order, n, degree, ℓ , and azimuthal number, m, of the underlying spherical harmonic, $Y_{n,\ell}^m$. While the most reliable method for pulsation mode identification is to analyze the line-profile variation using high-resolution spectroscopy (Mathias et al. 1997), our analysis remains limited to photometric data. Nonetheless, we apply three methods of mode identification based on photometry.

3.6.1. Mode identification based on the pulsation constant Q

The pulsation constant, Q, takes the unique value for any given pulsation frequency and can be used for mode identification (Breger & Bregman 1975). It is defined by $Q = P \sqrt{\bar{\rho}/\bar{\rho_{\odot}}}$, with P being the pulsation period and $\bar{\rho}$ and $\bar{\rho_{\odot}}$ the mean densities of the star and the Sun. Two important Q-values are 0.033 d and 0.026 d; they correspond to the fundamental and first radial overtone expected in δ Scuti stars. Expressing the densities as a function of the radius and eliminating the radius via the luminosity, the expression for Q can be recast (Breger 1990):

$$\log\left(\frac{Q}{P}\right) = 0.5\log(g) + 0.1\ M_{\text{bol}} + \log(T_{\text{eff}}) - 6.456. \tag{7}$$

Taking into account the uncertainties in the stellar parameters, Breger (1990) estimate the uncertainty in the *Q*-value to be 18%. For WASP-33A, we adopted $\log(g) = 4.3 \pm 0.2$, $d = 116 \pm 16$ pc, and $T_{\rm eff} = 7430 \pm 100$ K (Collier Cameron et al. 2010). We derived the absolute bolometric brightness $M_{\rm bol} = 2.85 \pm 0.07$

Table 6. Q values and errors for the eight frequencies found in our data.

ν (c/d)	Q(d)
$\nu_1 = 20.1621 \pm 0.0023$	$Q_1 = 0.035 \pm 0.008$
$v_2 = 21.0606 \pm 0.0023$	$Q_2 = 0.033 \pm 0.007$
$v_3 = 9.8436 \pm 0.0023$	$Q_3 = 0.071 \pm 0.016$
$v_4 = 24.8835 \pm 0.0017$	$Q_4 = 0.028 \pm 0.006$
$v_5 = 20.5353 \pm 0.0013$	$Q_5 = 0.034 \pm 0.007$
$v_6 = 34.1252 \pm 0.0027$	$Q_6 = 0.021 \pm 0.005$
$v_7 = 8.3084 \pm 0.0025$	$Q_7 = 0.085 \pm 0.019$
$v_8 = 10.8249 \pm 0.0030$	$Q_8 = 0.065 \pm 0.015$

using the expression $M_{\rm bol} = 42.36-5\log{(R/R_{\odot})} - 10\log{(T_{\rm eff})}$ (Allen 1973). For the eight pulsation frequencies in our pulsation model, we derive the Q-values listed in Table 6; errors have been estimated by error propagation.

Comparing the Q-values to the ones expected in δ Scuti stars (Breger 1998, and references therein), we find that ν_1 , ν_2 , ν_4 , ν_5 , and ν_6 are within the range of radial oscillations. Any further mode identification is not possible via Q-values.

To illustrate the difficulty of assigning modes accurately only by means of the pulsation constant, Q, we compare our most accurate Q_6 -value with the ones theoretically predicted by Fitch (1981). The model that best matches the WASP-33A parameters is labeled "1.5M21". Within errors, the following modes correspond to Q_6 : first, second, and third harmonic (Table 2A, radial modes); p_2 and p_3 (Table 2B, $\ell=1$ modes); p_1 , p_2 , and p_3 (Table 2C, $\ell=2$ modes); and p_1 , p_2 , and p_3 (Table 2D, $\ell=3$ modes). Therefore the only conclusive result is that Q_6 corresponds to a p-mode, which is expected for a high-frequency pulsation.

3.6.2. The empirical period-luminosity-color relation

Empirical period-luminosity-color (P-L-C) relations have been studied, e.g., by Petersen & Hog (1998), López de Coca et al. (1990), and King (1991), among many others.

Stellingwerf (1979) derive a theoretical P-L-C relation

$$\log P = -0.29 M_{\text{bol}} - 3.23 \log(T_{\text{eff}}) + \mathbb{C}, \tag{8}$$

where P is the period in days and \mathbb{C} is a constant equal to 11.96, 11.85, and 11.76 for the fundamental and first and second harmonics. Substituting our values for $M_{\rm bol}$ and $T_{\rm eff}$, yields $\nu_{0,\rm S}=23.43$ c/d, $\nu_{1,\rm S}=30.18$ c/d, and $\nu_{2,\rm S}=37.13$ c/d, i.e., periods that have not been identified within our data.

In an observational study, Gupta (1978) finds that a separate P-L-C relation for each pulsation mode provides a better agreement with the observations than a general one. The author derived the following empirical P-L-C relations for the fundamental mode, F (Eq. (9)), and the first, H1 (Eq. (10)), and second harmonic, H2 (Eq. (11)):

$$M_{\text{bol}\pm0.20} = -2.83\log(P_{\text{o}}) - 11.07\log(T_{\text{eff}}) + 41.61 \tag{9}$$

$$M_{\text{bol}\pm0.14} = -3.57 \log(P_1) - 10.21 \log(T_{\text{eff}}) + 37.13$$
 (10)

$$M_{\text{bol}\pm0.24} = -2.45 \log(P_2) - 10.22 \log(T_{\text{eff}}) + 38.35.$$
 (11)

These relations predicts $\nu_{0,G} = 27.91$ c/d, $\nu_{1,G} = 29.41$ c/d, and $\nu_{2,G} = 45.47$ c/d, again not observed within our data. At least for WASP-33A's stellar parameters, the theoretical and observational relation seem to be mutually inconsistent.

3.6.3. Multicolor photometry

In δ Scuti stars, the photometric amplitude and phase of pulsations depend on the spectral band. The amplitude and phase of a given pulsation are determined by the local effective temperature and cross-section changes, which are defined by the pulsation mode. Therefore different modes lead to distinguishable modulations in flux. This allows mode identification to be carried out by means of multicolor photometry (Balona & Evers 1999; Daszyńska-Daszkiewicz et al. 2003; Dupret et al. 2003).

Frequency Analysis and Mode Identification for Astero Seismology (FAMIAS) is a collection of software tools for the analysis of photometric and spectroscopic time series data (Zima 2008). The photometry module uses the method of amplitude ratios and phase differences in different photometric passbands to identify the modes (Balona & Stobie 1979; Watson 1988). The determination of the ℓ -degrees is based on static plane-parallel models of stellar atmospheres and on linear non-adiabatic computations of stellar pulsations. To compute the theoretical photometric amplitudes and phases, FAMIAS applies the approach proposed by Daszyńska-Daszkiewicz et al. (2002).

FAMIAS requires the stellar parameters' effective temperature, $T_{\rm eff}$, surface gravity, $\log g$, and metallicity, [Fe/H], which we obtained from Collier Cameron et al. (2010). As additional input to FAMIAS, we obtained the pulsation frequency, the amplitude, and the phase for the Strömgren v and b bands using our STELLA and CAHA data; amplitude ratios and phase differences were obtained using PERIOD04. Figure 7 shows our results for the case of v_2 , v_4 , and v_5 . The pulsations seem to correspond to lower-order modes: $\ell = 0$, 1 for v_5 and v_4 , and $\ell = 2$, 3 for v_2 . A more detailed characterization of these modes results impossible in our analysis. For the remaining five frequencies in our model, no reliable information on the associated modes could be derived.

3.6.4. The effect of rotation

Any asteroseismological study of main-sequence δ Scuti stars is not completely fulfilled until stellar rotation is considered (Goupil et al. 2000, and references therein). The effects of rotation over the pulsation spectrum has been theoretically studied (e.g., Deupree 2011; Deupree et al. 2012), as well as observed (e.g., Breger et al. 1999b, 2005, 2012). The main effect of rotation is the splitting of the nonradial mode frequencies. If such splitting is observed, then the rotation rate of a star can be determined (Christensen-Dalsgaard & Berthomieu 1991).

From a purely geometrical argument, stellar rotation affects the observed frequencies. In an inertial frame, an observer finds that a frequency is split uniformly according to the azimuthal order *m*:

$$v_m = v_0 + \Omega m, \tag{12}$$

where Ω is the angular velocity of the star, ν_0 is the frequency of the pulsation in the frame rotating with the star, and m the azimuthal mode. Using this simplified version of mode splitting (see, e.g., Cowling & Newing 1949, for the contribution of Coriolis forces to the frequency splitting), we produced the following analysis: from the eight frequencies that conform to our pulsation model, we assume that one of them, $\nu_{j,m}$, is the product of mode splitting. Therefore knowing that the azimuthal order m is associated with $\nu_{j,m}$ and the rotational period of the star, we can determine ν_0 . With ν_0 , we can further calculate the values of the remaining $\nu_{j,m}$ s for a given ℓ degree ($|m| < \ell$), and compare them with our remaining model frequencies.

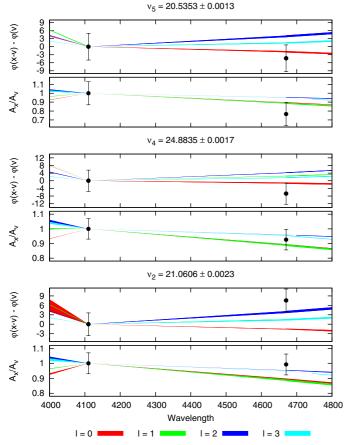


Fig. 7. Amplitude ratios and phase differences in degrees relative to the Strömgren v filter for WASP-33, resulting from the three available nights obtained at Calar Alto. The filled curves indicate the uncertainty of the theoretical prediction due to observational errors in $T_{\rm eff}$ and $\log g$.

Although this approach might sound straight forward, there is no knowledge of the rotational period of WASP-33A. We therefore assumed that WASP-33A's $v\sin(i)$ is coincident with the equatorial velocity. Furthermore, our attempt to identify the nature of the observed frequencies did not produce any substantial results. Consequently, for our most accurately identified v_2 frequency ($\ell=2,3$) we assumed all possible m values and found, through this reasoning, that the remaining observed frequencies were not the product of mode splitting.

Any further study would require, for instance, a complete mode identification and the knowledge of the rotational period of the star. The complexity around mode identification clearly indicates that Q values, P-L-C relations, rough period ratios, and even poor mode identification via multicolor photometry cannot be used for mode identification without further evidence.

4. Primary transit analysis

Our data comprise 19 primary transit observations. Table 7 lists, among others, the date and site of observation, the filter, and a code indicating the transit coverage of the observation. To determine the orbital parameters, we focus on the eight primary-transit light curves providing complete temporal coverage (TC = OIBEO in Table 7).

Moya et al. (2011) report on the detection of an optical companion about \sim 2" from WASP-33, which could affect our observations through third-light contamination. Based on the color information (J_c , H_c , K_c , and FII filters), the

Table 7. Summary of our primary transit observations and modeling parameters.

Fa Date Obs. TCb β1 β2 B 2011 Sept. 24 OLT - IBEO 2.61 2.15 2012 Aug. 23 OLT OIB - 3.14 2.47 2012 Oct. 11 OLT OIBE - 4.55 4.15 2012 Oct. 16 OLT OIBEO 4.41 3.87 V 2011 Oct. 10 OADM - IBE - 3.57 2.61 2011 Oct. 21 OAPS - IBEO 1.97 1.33 2012 Jan. 1 RNAV OIBEO 2.24 2.07 2012 Jan. 12 RNAV OIBEO 1.93 1.42 2012 Sept. 24 OADM - BEO 4.53 3.84 R 2010 Aug. 26 RNAV OIBEO 1.23 1.15 2010 Oct. 20 RNAV OIBEO 1.23 1.15	D
2012 Aug. 23 OLT OIB - 3.14 2.47	ν
2012 Oct. 11 OLT OIBE – 4.55 4.15 2012 Oct. 16 OLT OIBEO 4.41 3.87 V 2011 Oct. 10 OADM - IBE – 3.57 2.61 2011 Oct. 21 OAPS - IBEO 1.97 1.33 2012 Jan. 1 RNAV OIBEO 2.24 2.07 2012 Jan. 12 RNAV OIBEO 1.93 1.42 2012 Sept. 24 OADM - BEO 4.53 3.84 R 2010 Aug. 26 RNAV OIBEO 1.67 1.07 2010 Oct. 20 RNAV OIBEO 1.23 1.15	0
2012 Oct. 16 OLT OIBEO 4.41 3.87 V 2011 Oct. 10 OADM - IBE - 3.57 2.61 2011 Oct. 21 OAPS - IBEO 1.97 1.33 2012 Jan. 1 RNAV OIBEO 2.24 2.07 2012 Jan. 12 RNAV OIBEO 1.93 1.42 2012 Sept. 24 OADM - BEO 4.53 3.84 R 2010 Aug. 26 RNAV OIBEO 1.67 1.07 2010 Oct. 20 RNAV OIBEO 1.23 1.15	1
V 2011 Oct. 10 OADM - IBE - 3.57 2.61 2011 Oct. 21 OAPS - IBEO 1.97 1.33 2012 Jan. 1 RNAV OIBEO 2.24 2.07 2012 Jan. 12 RNAV OIBEO 1.93 1.42 2012 Sept. 24 OADM - BEO 4.53 3.84 R 2010 Aug. 26 RNAV OIBEO 1.67 1.07 2010 Oct. 20 RNAV OIBEO 1.23 1.15	1
2011 Oct. 21 OAPS — IBEO 1.97 1.33 2012 Jan. 1 RNAV OIBEO 2.24 2.07 2012 Jan. 12 RNAV OIBEO 1.93 1.42 2012 Sept. 24 OADM — BEO 4.53 3.84 R 2010 Aug. 26 RNAV OIBEO 1.67 1.07 2010 Oct. 20 RNAV OIBEO 1.23 1.15	1
2012 Jan. 1 RNAV OIBEO 2.24 2.07 2012 Jan. 12 RNAV OIBEO 1.93 1.42 2012 Sept. 24 OADM - BEO 4.53 3.84 R 2010 Aug. 26 RNAV OIBEO 1.67 1.07 2010 Oct. 20 RNAV OIBEO 1.23 1.15	0
2012 Jan. 12 RNAV OIBEO 1.93 1.42 2012 Sept. 24 OADM - BEO 4.53 3.84 R 2010 Aug. 26 RNAV OIBEO 1.67 1.07 2010 Oct. 20 RNAV OIBEO 1.23 1.15	0
Z012 Sept. 24 OADM - BEO 4.53 3.84 R 2010 Aug. 26 RNAV OIBEO 1.67 1.07 2010 Oct. 20 RNAV OIBEO 1.23 1.15	0
R 2010 Aug. 26 RNAV OIBEO 1.67 1.07 2010 Oct. 20 RNAV OIBEO 1.23 1.15	1
2010 Oct. 20 RNAV OIBEO 1.23 1.15	2
	0
**************************************	1
2010 Nov. 6 OLT OIBEO 3.24 2.85	2
2011 Oct. 5 RNAV OIBEO 2.05 1.63	0
2011 Nov. 23 RNAV OIBEO 1.95 1.52	0
v 2011 Oct. 22 CAHA – IBEO 1.41 1.18	0
2011 Nov. 2 STELLA OIBE – 2.36 1.95	0
b 2011 Oct. 22 CAHA – IBEO 1.49 1.37	0
2011 Nov. 2 STELLA OIBE – 2.48 2.27	0
y 2011 Oct. 22 CAHA – IBEO 1.52 1.06	0

Notes. Columns show the filter F, the date, the observatory where the transit was acquired, the transit coverage (TC), two coefficients, $\beta_{1,2}$, that account for correlated noise (Sect. 4.1), and the degree of the polynomial used for light curve normalization. ^(a) The filter B, V, and R for Johnson-Cousins and v, b, and y for Strömgren filters. ^(b) The letter code that specifies the transit coverage goes as follows: O for "out of transit before ingress", I for "ingress", B for "bottom", E for "egress", and O for "out of transit, after egress".

authors speculate that it might be a physical companion of WASP-33, for which they estimate an effective temperature of $T_{\rm eff} = 3050 \pm 250$ K. As the third-light contribution provided by such an object is $\lesssim 4 \times 10^{-4}$ in all used filters, it can be neglected in our analysis.

Orbiting a fast rotator in a quasi polar orbit (projected spinorbit misalignment $\lambda \sim 255^\circ$, Collier Cameron et al. 2010), the transit's light curves may be affected by gravity darkening, which manifests in a latitudinal dependence of the stellar effective temperature (von Zeipel 1924). As the rotational period of WASP-33 is unknown, we estimate it using $v \sin(i) \sim 90 \text{ km s}^{-1}$ and the stellar radius $R_s \sim 1.44 R_\odot$ (Collier Cameron et al. 2010). Close to the system geometry, we estimate the polar-to-equatorial temperature ratio. Using a gravity-darkening exponent of $\beta = 0.18$ (Claret 1998), for g the magnitude of the local effective surface gravity, and g the gravity-darkening exponent, following Maeder (2009):

$$T = T_{\text{pole}} \frac{g^{\beta}}{g^{\beta}_{\text{pole}}},\tag{13}$$

we estimate that the polar temperature of WASP-33 is $\approx 2.2\%$ higher than the equatorial temperature. This is too small to reproduce the observed transit depth wavelength-dependent variation. Further, using the primary transit code of Barnes (2009), which is adequate for fast rotators, we determine that the differences in the transit shape observed in the blue and red bands caused by gravity darkening are on the order of 0.06% and therefore negligible in our analysis. Therefore the *occultquad* routine provided by Mandel & Agol (2002) is adequate for our transit modeling.

4.1. Photometric noise

Often, the scatter in the light curve is used as a noise estimate. If, however, correlated noise is present, this method may considerably underestimate the impact of the scatter on the parameter estimates. The effect of correlated noise on transit modeling has been studied by several authors, e.g., Carter & Winn (2009); Pont et al. (2006).

While we have identified the significant pulsations in Sect. 3, our analysis has also shown that there is an unknown number of weak pulsations that we cannot account for in our modeling. The unaccounted pulsations will manifest in time-correlated noise in the transit analysis. Therefore a treatment of time-correlated noise is important in the transit modeling.

To quantify the amplitude of time-correlated noise in our data, we applied the "time-averaging method" proposed by Pont et al. (2006), which is based on the comparison of the variance of binned and unbinned residuals. To obtain the residuals, we normalized the transit light curves by fitting a polynomial to the out-of-transit data and then subtracted a preliminary transit model. We verified that the results only slightly depend on the details of the normalization and transit model.

Subsequently, the residual light curves were divided into M bins of equal duration. Each bin contains N data points. As our data are not always equally spaced, we applied a mean value for the number of data points per bin. In the absence of red noise, the expectation value of the variance of the unbinned residuals, σ_1 , is related to the variance of the binned residuals, σ_N , according to (cf. Carter & Winn 2009, Eq. (36))

$$\sigma_N = \sigma_1 \sqrt{\frac{M}{N(M-1)}}. (14)$$

This may now be compared with a variance estimate, σ'_N , derived from the binned residuals

$$\sigma'_{N} = \sqrt{\frac{1}{M} \sum_{i=1}^{M} (\overline{\mu} - \mu_{i})^{2}} \text{ with } \overline{\mu} = \frac{1}{M} \sum_{j=1}^{M} \mu_{j}.$$
 (15)

If correlated noise is present, then σ'_N will differ from σ_N by a factor β_N , which estimates the strength of correlated noise. A proper estimator, β , may be found by averaging β_N over a range Δn corresponding to the most relevant timescale. To account for the correlated noise in a conventional white-noise analysis, the individual photometric errors are enlarged by a factor of β . If there is no prior information, this leaves the parameter estimates unaffected and enlarges the errors by a factor β .

In the case of our transit analysis, the relevant timescale is the duration of ingress or egress, which is ~16 min for WASP-33. In Table 7 we show the resulting β factors. In a first step, we deliberately ignored our results derived in Sect. 3 and treated the light curves as if we had no knowledge on the pulsations. In the thus derived β_1 values, all pulsations show up as correlated noise. In a second step, we subtracted the pulsation model derived in Sect. 3 and determined the β_2 values. Taking into account the pulsation model always yields a better, that is, smaller β factor, indicating that the model accounts for a substantial fraction of the time correlation.

4.2. Polynomial order of transit light curve normalization

In transit modeling, normalization of the light curves is crucial. To normalize the transit light curves, we fit polynomials with an order between zero and four to the out-of-transit data and

Table 8. Linear (u_1) and quadratic (u_2) limb-darkening coefficients.

Filter	u_1	u_2
J-C B	0.3561	0.3625
J-CV	0.2725	0.3535
J-CR	0.1954	0.3511
Str v	0.3828	0.3678
Str b	0.3612	0.3512
Str y	0.2795	0.3533

determine the order, k, that minimizes the Bayesian information criterion, BIC = $\chi^2 + k \ln N$.

The order of the resulting optimal polynomial is listed in Table 7. According to our modeling, a constant or linear model is sufficient to normalize the transit in all but two cases, where a quadratic normalization is required. Finally, we visually inspected the resulting light curves to ensure a proper normalization.

4.3. Transit modeling

Collier Cameron et al. (2010) detect the planetary "shadow" of WASP-33b in the line profile of WASP-33A during transit. Their spectroscopic time series analysis reveals that the planet traverses the stellar disk at an inclination angle incompatible with 90°. As the inclination is less affected by parameter correlations in the spectroscopic analysis, we impose a Gaussian prior on the inclination. In particular, we use the value of $i = 87.67 \pm 1.8$ deg obtained by Collier Cameron et al. (2010) in their combined photometric and spectroscopic analysis.

In our analysis, we fixed the linear and quadratic limb-darkening coefficients, u_1 and u_2 , to the values listed in Table 8. We obtained these values, which consider the stellar parameters $\log(g) = 4.5$, $T_{\rm eff} = 7500$, and $[{\rm Fe/H}] = 0.1$, from Claret & Bloemen (2011). Therefore we are left with the following free parameters: the mid-transit time, $T_{\rm o}$, the orbital period, Per, the semimajor axis in stellar radii, $a/R_{\rm s}$, the orbital inclination, i, and the planet-to-star radius ratio, $p = R_{\rm p}/R_{\rm s}$. As previously mentioned, we used complete primary transits only to determine the best-fit orbital parameters.

To obtain the parameter estimates, their errors, and mutual dependence, we sample from the posterior probability distribution using an Markov chain Monte Carlo (MCMC) approach. For the parameters a, p, Per, and T_0 , we defined uniform priors covering a reasonable range. Errors are given as 68.3% highest probability density (HPD) credibility intervals. To carry out MCMC sampling, we used Python routines of PyAstronomy¹¹, which provide a convenient interface to fit and sample algorithms implemented in the PyMC (Patil et al. 2010) and SciPy (Jones et al. 2001) packages.

In a first attempt to fit the transits, we ignore the pulsations and fit only the transit light curves. In our approach, the complete transit light curves are fitted simultaneously using the model of Mandel & Agol (2002). We note that we also fitted the coefficients of the normalizing polynomial, whose degree remains, however, fixed to that listed in Table 7. Our best-fit solutions, which are obtained after 5×10^5 iterations, are given in Table 9.

In a second attempt, we combine the primary transit model with the pulsation model with frequencies and amplitudes fixed to the values listed in Table 4. In Sect. 3.5, we demonstrate that

¹¹ http://www.hs.uni-hamburg.de/DE/Ins/Per/Czesla/
PyA/PyA/index.html

Table 9. Orbital parameters of WASP-33.

Parameter	Value				
Collier Cameron et al. (2010)					
$a(R_{\rm s})$	3.79 ± 0.02				
i (°)	87.67 ± 1.81				
$p\left(R_{\rm p}/R_{\rm s}\right)$	0.1066 ± 0.0009				
Per (days)	$1.2198669 \pm 1.2 \times 10^{-6}$				
Kov	Kovács et al. (2013)				
$a(R_{\rm s})$	3.69 ± 0.01				
i (°)	86.2 ± 0.2				
$p(R_{\rm p}/R_{\rm s})$	0.1143 ± 0.0002				
Transit fit ignoring pulsations					
$a(R_{\rm s})$	3.69 ± 0.04				
i (°)	88.17 ± 1.53				
$p\left(R_{\rm p}/R_{\rm s}\right)$	0.1052 ± 0.0008				
Per (days)	$1.2198667 \pm 1.5 \times 10^{-6}$				
$T_{\rm o}~({\rm BJD_{TDB}})$	2455507.5225 ± 0.0004				
Transit fit a	Transit fit accounting for pulsations				
$a(R_s)$	3.68 ± 0.03				
<i>i</i> (°)	87.90 ± 0.93				
$p\left(R_{\rm p}/R_{\rm s}\right)$	0.1046 ± 0.0006				
Per (days)	$1.2198675 \pm 1.1 \times 10^{-6}$				
$T_{\rm o}~({\rm BJD_{TDB}})$	2455507.5222 ± 0.0003				

Notes. Parameters obtained by Collier Cameron et al. (2010) and Kovács et al. (2013); our best-fit results were obtained from primary transit modeling: first, ignoring pulsations, and second, taking into account pulsations.

there is a temporal evolution in the phases. Therefore the phases have been considered free parameters in our modeling. However, we did not allow them to take arbitrary values but restricted the allowed range to the limiting cases derived in Sect. 3.5. For instance, the phase of the first frequency, ν_1 , could not deviate by more than 0.1 cycle from the mean value (cf. Fig. 6).

The results are presented in the lower part of Table 9; in Fig. 8 we show the 19 primary transit light curves and the best-fit model. Interestingly, the parameters derived using this more elaborate approach are consistent with those obtained ignoring the pulsations. Taking into account the pulsation model does, however, improve the uncertainties in the parameter estimates with respect to the regular primary transit-fitting approach.

The values derived in our analysis are broadly consistent with those derived previously by Collier Cameron et al. (2010) and Kovács et al. (2013). While we find a slightly smaller semimajor axis than Collier Cameron et al. (2010)'s, the planet-to-star radius ratio and the inclination are compatible. Kovács et al. (2013) find an 8–10% larger radius ratio and a slightly lower inclination. A homogeneous study of all the primary transits available in the bibliography escapes the purpose of this work. However, we believe that the small differences in the orbital parameters observed by different authors might be the product of an inadequate normalization of the primary transits or an insufficiently correlated noise treatment.

4.3.1. Impact of the pulsations on the transit fits

To better understand the effect of the pulsations on the transit fits, we fit the 19 primary transits individually and study the behavior of a, i, and p. We carry out the fit one more time first ignoring the pulsations and then taking them into account via our pulsation model. During the fit, the ephemeris were fixed to the corresponding values in Table 9. The outcomes are based

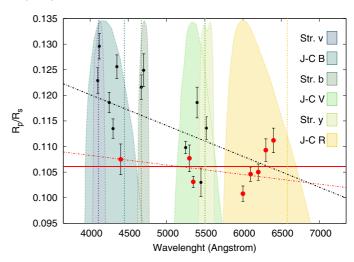


Fig. 9. Planet-to-star ratio $\hat{p} = R_{\rm p}/R_{\rm s}$ obtained from complete light curves (red points) and incomplete ones (black points) when the pulsations have been accounted for in the model fitting. Vertical color-dashed lines indicate the central wavelength of each filter. The dashed/dotted black line shows the best-fitting linear model to the 19 $R_{\rm p}/R_{\rm s}$ data points. In red and considering complete primary transits only, the dashed/dotted line shows the low-significant wavelength-dependent trend, while the continuous line accounts for the mean radius ratio.

on 5×10^5 iterations of the MCMC sampler; they are given in Table 10.

To study the impact of the pulsation model on the individual parameters, we scrutinized the ratio of the derived values. In particular, we focused on the eight complete transits whose parameters can be determined most reliably. For the ratios of values determined with pulsations considered in the model (wp) and neglected pulsations (pn), we obtained $a_{wp}/a_{pn} = 1.0 \pm 0.02$, $i_{wp}/i_{pn} = 1.000 \pm 0.001$, and $p_{wp}/p_{pn} = 0.99 \pm 0.03$. These numbers indicate that, on average, the parameter estimates remain unaffected by taking into account the stellar pulsations. Regarding individual fits, the expected deviation amounts to $0.08 R_s$ in the semimajor axis, 0.1° in the inclination, and 3 \times 10^{-3} in p. Clearly, the relative uncertainty is largest in the semimajor axis and the radius ratio, p. Also taking into account the transits with incomplete observational coverage, we obtain numbers that are comparable but with larger uncertainties. It is worth mentioning that the "outlying" orbital parameters, which are presented in Table 10, are the product of incomplete primary transit fitting. Therefore we believe that primary transit normalization might play a fundamental role in the determination of such parameters.

4.4. Wavelength dependence of the planet-to-star radius ratio

Our data comprise transit observations from the blue to the red filter. To check whether a dependence of the planet-to-star radius ratio on the wavelength can be identified, we fixed all parameters but the radius ratio to the values listed in Table 9 and fitted only the radius ratio, \hat{p} , for each individual transit. The resulting \hat{p} values, which are based on the pulsation-corrected light curves, are listed in the last column of Table 9. We verified that we obtain comparable results if the pulsations are not considered.

Figure 9 shows \hat{p} as a function of wavelength. The red points mark the transits with full observational coverage, while the black points were derived from transits with incomplete temporal coverage (see Table 7). In the fit, we used the central

wavelength of the filters, which are indicated by vertical, dashed lines in the plot. To produce a less crowded figure, the data points are artificially shifted from the central wavelength.

The values for \hat{p} , which are derived from the complete transits, show a wavelength-dependent trend, but only with marginal significance. These values are also consistent with a constant radius ratio of $R_{\rm P}/R_{\rm S}=0.1061\pm0.0031$, in concordance with the mean value and the standard deviation of the eight data points. Both the linear and constant models produce comparable $\chi^2_{\rm red}$ values of 1.9 and 1.8, respectively. Hence, it is unclear which model best fits the data. If all transit observations are taken into account, the data indicate a decrease of 0.65%/1000 Å in the planet-to-star radius ratio with wavelength. Formally, the correlation coefficient between radius-ratio and wavelengths amounts to $r\sim0.7$. We caution, however, that the observed trend may be feigned by an inappropriate transit normalization because many transits lack full observational coverage.

Kovács et al. (2013) also notice a substantial difference in the transit depth derived from an observation in $H\alpha$ and a simultaneously obtained J-band light curve (their Fig. 9). In particular, the J-band transit is shallower, which would be consistent with the wavelength trend. However, the signature observed through the $H\alpha$ line may significantly differ from those observed in broadband filters because the $H\alpha$ line is affected by strong chromospheric contributions.

5. Discussion

5.1. The stellar pulsation spectrum

The pulsation spectrum of WASP-33 has been studied by several authors, who report a wealth of frequencies (see Table 11). All studies find pulsations with amplitudes on the order of 1 mmag, which is compatible with our results.

Deming et al. (2012) observed WASP-33 during two secondary transits in the K_s band using the 2.1 m telescope at Kitt Peak National Observatory and for another two nights using the *Spitzer* telescope. All observations were performed during secondary transits. Their frequency analysis was carried out for individual nights. Their first three frequencies (21.1, 20.2, and 9.8 c/d) are compatible with our v_1 , v_2 , and v_3 (see Table 4). Also, de Mooij et al. (2013) observed secondary transits of WASP-33b in the K_s band for two nights, each lasting ~5 h. Although the frequencies they report are within the range of values we find, the values are numerically inconsistent with our results.

Herrero et al. (2011) observed for nine nights mainly during primary transit using Johnson-Cousins R filter. They report pulsation frequencies, which might correspond to our v_1 . Smith et al. (2011) carried out observations during one secondary transit using an S[III] narrow-band filter centered at 9077 Å. The observations were performed for almost nine hours during a single night lacking photometric conditions. Among the pulsation frequencies they find, 21.6 ± 0.6 c/d and 34.3 ± 0.4 c/d likely correspond to our v_2 and v_6 .

The most extensive pulsation analysis is carried out by Kovács et al. (2013). It is based on four photometric datasets, including that of Herrero et al. (2011). Kovács et al. (2013) report two frequencies that are compatible with our v_1 and v_2 . The ~15.2 c/s frequency does not show up in our analysis.

Among the different data sets, pulsations with frequencies around ~ 21 c/d and ~ 20 c/d that correspond to our most pronounced frequencies ν_1 and ν_2 consistently occur. Other pulsation frequencies are only found in some cases. We note,

Table 11. Evolution of the reported frequencies and amplitudes for WASP-33's pulsation spectrum (from top to bottom).

Frequency	Amplitude
(c/d)	(mmag)
Herrero et a	ol (2011)
21.004 ± 0.004	0.98 ± 0.05
21.311 ± 0.004	~0.86
Smith et al	
26.9 ± 0.4	1.479 ± 0.069
18.8 ± 0.6	0.567 ± 0.134
34.3 ± 0.4	0.766 ± 0.115
21.6 ± 0.6	0.605 ± 0.115 0.605 ± 0.105
Deming et a	
~21.1	~1.3 ~2.3
~20.2 ~9.8	~2.3 ~1.6
~9.8	~1.0
~20.6	~2.1 ~2.1
de Mooij et a	al. $(2013)^a$
22.5 ± 0.1	0.95 ± 0.04^{N1}
33.3^{N1}	0.41 ± 0.04
27.3 ± 0.2	0.56 ± 0.06^{N2}
33.2^{N2}	0.17 ± 0.05
22.0^{N1}	0.11 ± 0.06
17.1^{N1}	0.13 ± 0.06
Kovács et al	$(2013)^b$
15.21643 ± 0.00004	0.758 ± 0.085^{HN}
20.16229 ± 0.00004	0.733 ± 0.080^{HN}
21.06339 ± 0.00004	0.719 ± 0.078^{HN}
15.21517 ± 0.00001	$0.477 \pm 0.054^{H+F}$
20.16230 ± 0.00001	$0.739 \pm 0.053^{H+F}$
21.06346 ± 0.00001	$0.728 \pm 0.049^{H+F}$
This w	ork
20.1621 ± 0.0023	1.03 ± 0.03
21.0606 ± 0.0023	1.01 ± 0.03
9.8436 ± 0.0023	0.86 ± 0.03
24.8835 ± 0.0017	0.45 ± 0.03
20.5353 ± 0.0013	0.77 ± 0.03
34.1252 ± 0.0027	0.53 ± 0.03
8.3084 ± 0.0025	0.68 ± 0.03
10.8249 ± 0.0030	0.69 ± 0.03

Notes. ^(a) Following the nomenclature of de Mooij et al. (2013), the superscripts N1 and N2 refer to "Night I" and "Night II". ^(b) The superscript HN stands for "HATNet" and H+F for "HATNet+FUP".

however, that the data were acquired in different spectral bands, mostly in the infrared, where pulsations are expected to be lower in amplitude. The residuals after subtracting our pulsation model clearly indicate the presence of further low-amplitude pulsations, which might correspond to those found in previous studies. Additionally, our phase shift analysis in Sect. 3.5 has shown that the pulsation spectrum might be intrinsically variable. The amplitudes of the pulsations found by the various studies are all on the order of 1 mmag, which is compatible with our results. As the amplitudes are intrinsically small and furthermore wavelength dependent, we refrain from a detailed comparison of the derived numbers.

In our analysis, we identify eight significant pulsation frequencies. Although the pulsation spectrum is probably much more complex than that, the amplitudes of the pulsations are intrinsically low and more data is required to characterize the components in further detail. We show that the pulsation phases

vary in time with a gradient, dp, of up to $|dp| \le 2 \times 10^{-3} \, d^{-1}$, assuming a linear evolution. This suggests that the amplitudes and frequencies also show temporal variability. However, our data do not allow to verify this.

We find that most of the detected frequencies are likely associated with low-order p-modes. We attempt to further identify the modes using Q values, empirical P-L-C relations, and amplitudes and phases of multicolor photometry. However, we find the detected frequencies to be largely incompatible with all these relations. We argue that this is not uncommon (e.g. Breger et al. 2005).

5.2. Transit modeling

In their analysis of photometric follow-up data, Kovács et al. (2013) find a persistent "hump" in the residuals obtained after subtracting the transit model shortly after mid-transit time. While our residuals clearly show unaccounted pulsations, we do not see any such hump recurring at the same phase. Therefore we find no evidence for a persistent structure on the stellar surface like, e.g., a spot belt as suggested by Kovács et al. (2013).

Although our transit modeling is consistent with a constant star-to-planet radius ratio with respect to wavelength, there may be a slight trend indicating a decrease in the radius ratio as the wavelength increases. Although this would be compatible with the results of Kovács et al. (2013), who find that the radius ratio in the J band is smaller than that observed in $H\alpha$, we caution that the formation of the $H\alpha$ line may be different from that observed in broadband filters – a caveat already mentioned by Kovács et al. (2013). Based on the currently available data, we conclude that a constant planet-to-star radius ratio seems most likely.

5.3. Star-planet interaction

Collier Cameron et al. (2010) report on a nonradial pulsation at a frequency of \sim 4 c/d, which might be tidally induced by the planet. Unfortunately, this frequency lies outside the sensitivity range of our analysis.

In close binary systems, tidal interaction affects stellar oscillations (e.g., Cowling 1941; Savonije & Witte 2002; Willems 2003). In particular, Hambleton et al. (2013) studied a short-period binary system that presents δ Scuti pulsations and tidally excited modes. In addition to the already known commensurability between the pulsation frequencies and the orbital period of the system, the authors found that the spacing between the detected p-modes was an integer multiple of the system's orbital frequency. Although it is clear that the nature of WASP-33 does not resemble a short-period eccentric binary system, in order to analyze star-planet interaction we search for commensurability of the detected pulsation frequencies with the planetary orbital period and investigated the spacings between them.

As the exact rotation period of the host star WASP-33A remains unknown, it is not entirely clear how exactly the planet affects the stellar surface in the frame of the star; in particular, the period at which the planet affects the same surface element is unknown for the larger fraction of the stellar surface. However, as the planet orbits in a highly tilted, nearly polar orbit, the stellar poles experience a periodic force with a period identical to once and twice the planetary orbital period. When the planet crosses a pole, the effective gravity on both poles is lowered due to the planetary gravity and orbital motion; the effect is, however, not the same on both poles.

Using our best-fit orbital period of 1.2198675 d, we express the pulsation frequencies in terms of the orbital frequency of the planet. The result is presented in the last column of Table 4, where the ratios of the pulsations and the orbit are displayed. We expect the error in the pulsation frequencies to be considerably larger than those in the orbital period. The closest commensurability is found for the 9.8436 c/d, which corresponds to 12.008 times the orbital frequency.

To assess the significance of such a result, we carry out a Monte-Carlo simulation. In particular, we randomly generate eight frequencies between 8 and 34, i.e., in the approximate range of our detected pulsations. The frequencies are dropped out from a uniform distribution. We then calculate the associated ratio based on the orbital frequency and finally search for the best match. After 50 000 runs, we find that the cumulative probability distribution for the minimum distance from an integer frequency ratio is given by

$$F(d_{\min}) = 1 - e^{-\frac{d_{\min}}{0.060934}},$$

obtained after fitting our Monte-Carlo results with an exponential decay. Using this relation, we find that the probability of detecting at least one of the ratios as close or closer than 0.008 c/d to an integer ratio to be 12%. The ratio may indeed be an integer, considering the error in the frequency determination (see Sect. 3). Nevertheless, we note that our phase-shift analysis also revealed variability in the frequency spectrum, which we find hard to reconcile with a simple picture of tidally excited pulsations.

Additionally, we found that the spacing between the frequencies cannot be described by harmonics of the orbital period of the system. In fact, the best case scenario is given by v_5 and v_7 . Considering our best-fit orbital period, the departure from an integer number is ten times their estimated error.

Therefore we conclude that there is no evidence for a direct relation between any of our pulsation frequencies and the planetary orbital period.

6. Conclusions

In this work, we obtained and analyzed an extensive set of photometric data of the hottest known star hosting a hot Jupiter, WASP-33. The data cover both in- and out-of-transit phases and are used to study the pulsation spectrum and the primary transits.

In particular, our out-of-transit data provide ~ 3 times more temporal coverage than the Kovács et al. (2013) data set, which is the most extensive among those listed in Table 11. In addition, our data set is the only one that comprises dedicated out-of-transit photometric coverage to study the stellar pulsations in detail, along with multicolor and simultaneous observations to study the nature of the modes.

A comprehensive study of the pulsation spectrum of WASP-33 reveals, for the first time, eight significant frequencies. Additionally, some of the frequencies found seem to be consistent with previous reports. Along with the associated amplitudes and phases, we construct a pulsation model that we use to correct the primary transit light curves with the main goal of redetermining the orbital parameters by means of pulsation-clean data.

In our transit modeling, we find the system parameters broadly consistent with those reported by Collier Cameron et al. (2010) and Kovács et al. (2013). Interestingly, the derived parameter values are hardly affected by taking into account the pulsations in the modeling although the errors decrease. This

statement clearly depends on the total number of observed transits and the stability of the p-modes. Thus further observations of primary transit events of WASP-33 will be required to support or overrule this remark.

One possible explanation of the behavior of the orbital parameters with respect to the pulsations of the host could be that the associated amplitudes, at least in the high-frequency range that our studies focus on, are small in nature. Furthermore, our extensive primary transit observations, which we obtained in different filter bands, allow us to notice a decrease in the planet-tostar radius ratio with wavelength. This decrease has also been observed by other authors. Simultaneous multiband photometry of primary transits of WASP-33 will help to better constrain this dependency.

Considering that our work was produced using fully groundbased observations, we were able to provide an extensive study of the pulsation spectrum of this unique δ Scuti host star. This, in turn, has helped to better comprehend how much pulsations affect the determination of system parameters.

Acknowledgements. C. von Essen acknowledges funding by the DFG in the framework of RTG 1351, Dr. Andres Moya and Prof. Rafael Garrido for discussing issues on δ Scuti stars, and the anonymous referee for her/his improvements on the draft. E. H. and I. R. acknowledge financial support from the Spanish Ministry of Economy and Competitiveness (MINECO) and the "Fondo Europeo de Desarrollo Regional" (FEDER) through grant AYA2012-39612-C03-01. The Joan Oró Telescope (TJO) of the Montsec Astronomical Observatory (OAdM) is owned by the Generalitat de Catalunya and operated by the Institute for Space Studies of Catalonia (IEEC). We further thank Ramon Naves for operating the 0.3 m telescope at Observatori Montcabrer, Thomas Granzer for his support on STELLA observations and data reduction, and H. Hagen for his technical support at OLT.

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Table 1. Overview of observation nights for OLT and STELLA.

Date	Duration (h)	F^a	NoP^b	Airmass			
	OLT						
2010 November 6	5.9	R	803	$1.0 \to 1.2$			
2010 December 1	1.4	B	60	$1.5 \rightarrow 2.1$			
2011 February 3	5.2	R	869	$1.0 \rightarrow 1.4$			
2011 February 8	4.0	R	89	$1.0 \rightarrow 1.2$			
2011 February 21	5.9	R	1277	$1.0 \rightarrow 2.3$			
2011 February 22	5.3	R	769	$1.0 \rightarrow 2.3$			
2011 March 1	4.0	B	83	$1.0 \rightarrow 1.6$			
2011 March 3	3.6	B	176	$1.0 \rightarrow 1.6$			
2011 March 8	2.5	B	186	$1.2 \rightarrow 2.3$			
2011 March 11	4.6	B	296	$1.1 \rightarrow 2.4$			
2011 March 20	3.3	B	277	$1.2 \rightarrow 2.1$			
2011 August 31	5.5	B	541	$1.0 \rightarrow 2.4$			
2011 September 1	5.2	B	906	$1.0 \rightarrow 2.5$			
2011 September 23	4.1	B	632	$1.0 \rightarrow 2.9$			
2011 September 24	4.3	B	632	$1.0 \rightarrow 1.2$			
2011 September 30	5.8	B	754	$1.1 \rightarrow 2.0$			
2011 October 1	3.2	B	443	$1.4 \rightarrow 3.1$			
2011 October 2	3.5	B	439	$1.5 \rightarrow 2.2$			
2011 November 28	7.1	B	650	$1.0 \rightarrow 1.4$			
2011 November 30	7.6	B	1174	$1.0 \rightarrow 2.4$			
2012 August 23	5.2	B	545	$1.0 \rightarrow 2.3$			
2012 August 25	6.5	B	242	$1.0 \rightarrow 1.8$			
2012 August 28	3.9	B	243	$1.3 \rightarrow 2.8$			
2012 September 12	7.3	B	621	$1.0 \to 2.1$			
2012 October 8	6.6	B	758	$1.0 \rightarrow 2.4$			
2012 October 11	8.6	B	1160	$1.0 \rightarrow 2.0$			
2012 October 15	2.9	B	241	$1.3 \rightarrow 2.1$			
2012 October 16	4.9	B	667	$1.0 \rightarrow 1.5$			
2012 October 28	11.8	В	736	$1.0 \rightarrow 2.6$			
STELLA							
2011 October 25	11	v	78	$1.0 \rightarrow 2.9$			
	11	b	81				
2011 October 27	10.4	v	301	$1.0 \rightarrow 2.6$			
	10.4	b	144				
2011 October 28	10.1	v	322	$1.0 \rightarrow 2.5$			
	10.1	b	318				
2011 October 29	6.7	v	254	$1.0 \rightarrow 2.6$			
	5.4	b	194				
2011 November 2	8.5	v	326	$1.0 \rightarrow 2.2$			
	8.5	b	320				
2011 November 13	4.4	v	123	$1.0 \rightarrow 1.9$			
	4.4	b	116				
2011 November 14	10.4	v	357	$1.0 \rightarrow 2.7$			
	10.4	b	358				
2011 November 15	6.4	v	236	$1.0 \rightarrow 1.8$			
	4.7	b	203				
2011 November 26	4.8	v	175	$1.0 \rightarrow 1.5$			
	4.8	b	187				
				10 . 17			
2011 November 27	8.6	v	311	$1.0 \rightarrow 1.7$			

Notes. (a) Filter; (b) Number of photometric data points.

Table 10. Results of individual transit fits.

Date	F	$a(R_s)$	i (°)	p	p̂		
]	Pulsations ignore	ed			
2010 Aug. 26 *	R	3.88 ± 0.12	87.74 ± 1.77	0.0999 ± 0.0026	0.0988 ± 0.0027		
2010 Oct. 20 *	R	3.72 ± 0.06	88.05 ± 1.60	0.1006 ± 0.0016	0.1004 ± 0.0016		
2010 Nov. 6 *	R	3.69 ± 0.08	87.62 ± 1.78	0.1095 ± 0.0019	0.1071 ± 0.0023		
2011 Sept. 24	B	3.63 ± 0.06	88.07 ± 1.62	0.1239 ± 0.0021	0.1231 ± 0.0019		
2011 Oct. 5 *	R	3.50 ± 0.10	87.73 ± 1.76	0.1145 ± 0.0033	0.1131 ± 0.0032		
2011 Oct. 10	V	3.44 ± 0.17	87.76 ± 1.73	0.1052 ± 0.0035	0.1041 ± 0.0031		
2011 Oct. 21	V	3.48 ± 0.06	87.94 ± 1.69	0.1111 ± 0.0018	0.1101 ± 0.0017		
2011 Oct. 22	v	3.68 ± 0.10	87.72 ± 1.79	0.1241 ± 0.0028	0.1238 ± 0.0027		
2011 Oct. 22	b	3.56 ± 0.10	87.61 ± 1.82	0.1223 ± 0.0026	0.1217 ± 0.0025		
2011 Oct. 22	y	3.61 ± 0.11	87.74 ± 1.75	0.1137 ± 0.0033	0.1134 ± 0.0031		
2011 Nov. 2	v	3.92 ± 0.17	87.36 ± 1.87	0.1311 ± 0.0032	0.1287 ± 0.0032		
2011 Nov. 2	b	3.92 ± 0.21	87.34 ± 1.92	0.1271 ± 0.0035	0.1248 ± 0.0036		
2011 Nov. 23 *	R	3.56 ± 0.09	87.64 ± 1.80	0.1101 ± 0.0030	0.1088 ± 0.0030		
2012 Jan. 1st *	V	3.43 ± 0.09	87.73 ± 1.77	0.1113 ± 0.0033	0.1093 ± 0.0035		
2012 Jan. 12 *	V	3.70 ± 0.07	87.93 ± 1.69	0.1058 ± 0.0018	0.1061 ± 0.0017		
2012 Aug. 23	B	4.21 ± 0.18	86.94 ± 2.07	0.1174 ± 0.0024	0.1093 ± 0.0026		
2012 Sep. 24	V	3.83 ± 0.14	87.77 ± 1.75	0.1167 ± 0.0037	0.1176 ± 0.0032		
2012 Oct. 11	B	3.48 ± 0.12	87.69 ± 1.79	0.1115 ± 0.0032	0.1131 ± 0.0032		
2012 Oct. 16 *	В	3.63 ± 0.10	87.71 ± 1.76	0.1098 ± 0.0031	0.1097 ± 0.0031		
	Pulsations taken into account						
2010 Aug. 26 *	R	3.74 ± 0.07	87.80 ± 1.71	0.1011 ± 0.0015	0.1008 ± 0.0015		
2010 Oct. 20 *	R	3.73 ± 0.06	88.11 ± 1.59	0.1046 ± 0.0014	0.1046 ± 0.0014		
2010 Nov. 6 *	R	3.61 ± 0.07	87.83 ± 1.69	0.1050 ± 0.0016	0.1050 ± 0.0016		
2011 Sept. 24	B	3.64 ± 0.06	88.20 ± 1.57	0.1194 ± 0.0024	0.1186 ± 0.0020		
2011 Oct. 5 *	R	3.49 ± 0.09	87.62 ± 1.73	0.1143 ± 0.0024	0.1112 ± 0.0024		
2011 Oct. 10	V	3.59 ± 0.11	87.88 ± 1.70	0.1046 ± 0.0039	0.1030 ± 0.0028		
2011 Oct. 21	V	3.54 ± 0.06	87.79 ± 1.75	0.1103 ± 0.0013	0.1098 ± 0.0012		
2011 Oct. 22	v	3.49 ± 0.13	87.57 ± 1.84	0.1236 ± 0.0008	0.1229 ± 0.0025		
2011 Oct. 22	b	3.33 ± 0.10	87.56 ± 1.83	0.1235 ± 0.0024	0.1216 ± 0.0024		
2011 Oct. 22	y	3.52 ± 0.13	87.65 ± 1.77	0.1143 ± 0.0023	0.1136 ± 0.0022		
2011 Nov. 2	v	3.79 ± 0.17	87.03 ± 1.00	0.1212 ± 0.0016	0.1296 ± 0.0025		
2011 Nov. 2	b	3.90 ± 0.25	87.14 ± 1.96	0.1277 ± 0.0034	0.1249 ± 0.0032		
2011 Nov. 23 *	R	3.53 ± 0.08	87.64 ± 1.83	0.1126 ± 0.0023	0.1093 ± 0.0022		
2012 Jan. 1 st *	V	3.54 ± 0.09	87.76 ± 1.75	0.1082 ± 0.0026	0.1077 ± 0.0026		
2012 Jan. 12 *	V	3.73 ± 0.07	87.80 ± 1.69	0.1034 ± 0.0012	0.1031 ± 0.0011		
2012 Aug. 23	В	4.00 ± 0.12	87.29 ± 1.87	0.1157 ± 0.0019	0.1135 ± 0.0019		
2012 Sep. 24	V	3.89 ± 0.09	87.70 ± 1.75	0.1206 ± 0.0033	0.1186 ± 0.0030		
2012 Oct. 11	B	3.27 ± 0.09	87.42 ± 1.85	0.1245 ± 0.0023	0.1256 ± 0.0023		
2012 Oct. 16 *	В	3.64 ± 0.10	87.79 ± 1.71	0.1054 ± 0.0027	0.1075 ± 0.0030		

Notes. Complete primary transits are indicated with an *.

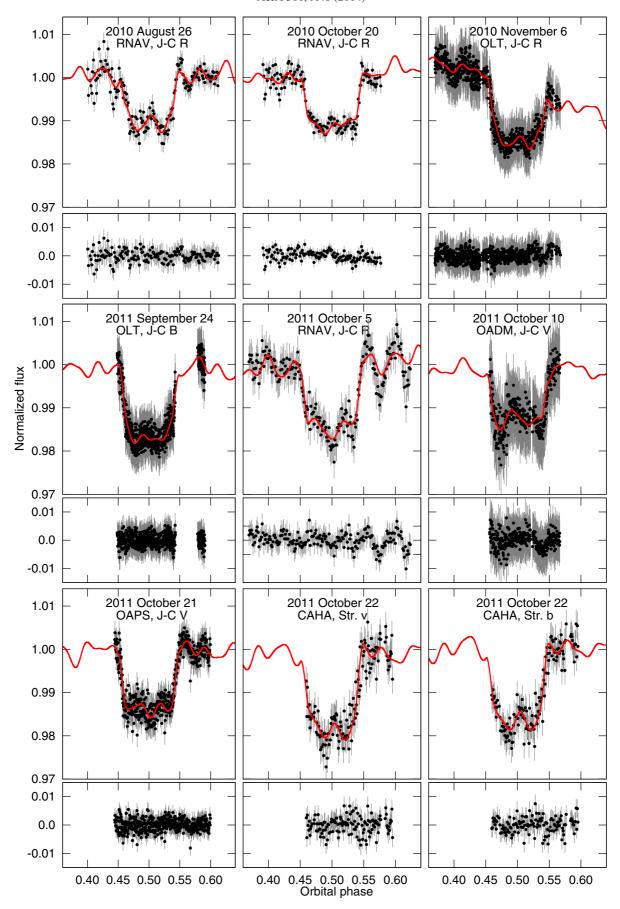


Fig. 8. *Top panels*: the 19 primary transits in black points, along with the photometric error bars accounting for correlated noise, cf. Sect. 4.1. Overplotted in continuous red line is the best-fitted primary transit model modulated by the host star pulsations and the low-order normalization polynomial. *Bottom panels*: residuals.

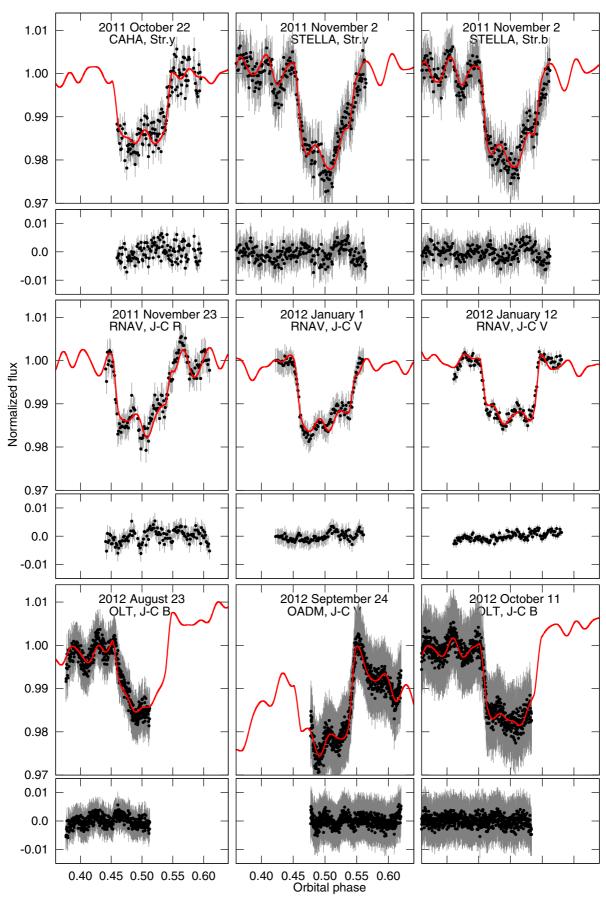


Fig. 8. continued.

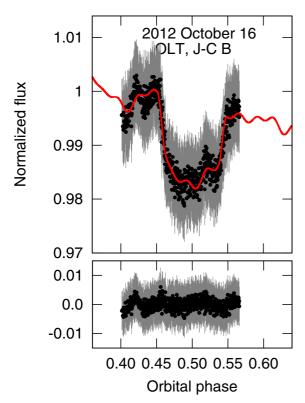


Fig. 8. continued.